

Section 9

Regression Discontinuity Designs

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GOV 2003

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Overview

- Logistics:
 - **Pset 8 (the last pset) released!:** Due at 11:59 pm (ET) on Nov 17
 - **November 12th 19th:** Submit a brief (no longer than 5 page) page memo of your main results, including tables, figures, and brief analysis. For methodological projects, this should include a description of the method and any analytical/simulation results. You will be required to give feedback on another group's project, which will be counted toward the overall grade based on attentiveness and usefulness of the feedback provided.
- Today's topics:
 - Sharp RD
 - Identification
 - Estimation
 - Diagnostics

Sharp RD

- Finding exogenous variation in the treatment assignment
 - RD: a **discontinuity** in treatment *assignment*
- Example: incumbency advantage in the U.S. House (Lee 2008)
 - “The overall causal impact of being the current incumbent party in a district on the votes obtained in the district’s election”
 - Treatment (D_i): being the current incumbent party
 - Forcing (X_i): margin of victory (at election t)
 - Outcome (Y_i): probability of winning (at election $t + 1$)
 - Sharp RD: $D_i = 1\{X_i \geq c\} \forall i$
- Estimand: local average treatment effect **at the cutoff**

$$\begin{aligned}\tau_{\text{srd}} &= \mathbb{E}[Y_i(1) - Y_i(0)|X_i = c] \\ &= \mathbb{E}[Y_i(1)|X_i = c] - \mathbb{E}[Y_i(0)|X_i = c]\end{aligned}$$

Sharp RD

- Identifying the effect at the cutoff with **continuity** of CEFs

$$\tau_{\text{srd}} = \underbrace{\lim_{x \downarrow c} \mathbb{E}[Y_i | X_i = x]}_{=\mathbb{E}[Y_i(1)|X_i=c]} - \underbrace{\lim_{x \uparrow c} \mathbb{E}[Y_i | X_i = x]}_{=\mathbb{E}[Y_i(0)|X_i=c]}$$

- Estimate the limit by the local linear regression

$$(\widehat{\alpha}_+, \widehat{\beta}_+) = \underset{i: X_i \geq c}{\text{argmin}} \sum \{Y_i - \alpha - \beta(X_i - c)\}^2 \underbrace{K\left(\frac{X_i - c}{h}\right)}_{\text{weights}}$$

- We take the estimated intercept: $\widehat{\alpha}_+ = \widehat{\mathbb{E}}[Y_i(1) | X_i = c]$
- Our point estimate is: $\widehat{\tau}_{\text{srd}} = \widehat{\alpha}_+ - \widehat{\alpha}_-$

Graphical Illustration

Identification: continuity v. local randomization

- Continuity assumption is not equivalent to local randomization

$$\{Y_i(1), Y_i(0)\} \perp\!\!\!\perp 1\{X_i > c\} \mid c_0 \leq X_i \leq c_1$$

- Stronger than continuity. Why?
- Estimation and its visualization?

Estimation and visualization

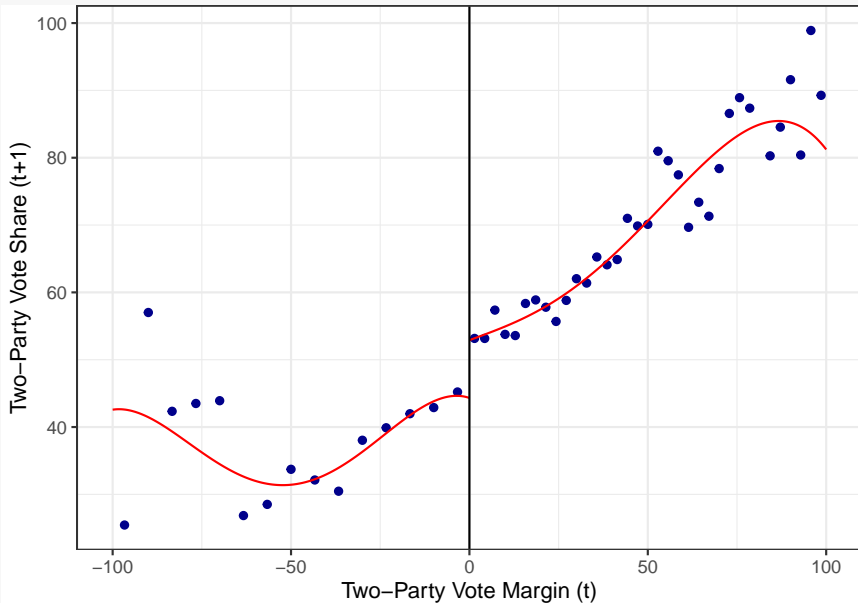
- Use `rdr` package (current standard)

1. Visualization: showing discontinuity at the cutoff

plot (data-driven Regression Discontinuity (RD) plots)

```
rdplot(y = vote, x = margin, kernel = "tri",  
       title = "", y.label = "Two-Party Vote Share (t+1)",  
       x.label = "Two-Party Vote Margin (t)")
```

RD Plot



Estimation and visualization

2. Estimation: fit one linear regression with the interaction between $(X_i - c)$ and D_i

$$\underset{(\alpha, \beta, \tau, \gamma)}{\operatorname{argmin}} \sum_{i: X_i \in [c-h, c+h]} \{Y_i - \alpha - \beta(X_i - c) - \tau D_i - \gamma(X_i - c)D_i\}^2$$

3. Optimal bandwidth, bias correction and robust standard errors
- Intuition:
 - find bandwidth that minimizes the estimation error
 - \leadsto we don't know the true bias and have to estimate it
 - \leadsto additional uncertainty
 - Calonico, Cattaneo, and Titiunik (CCT, 2014, Econometrica)

fit local linear regression

```
fit <- rdrobust(y = vote, x = margin, p = 1, kernel = "tri")
```

Results

| | estimate | se |
|----------------|----------|-------|
| Conventional | 7.414 | 1.459 |
| Bias-Corrected | 7.507 | 1.459 |
| Robust | 7.507 | 1.741 |

- Two types of point estimates:
 1. The standard local linear estimator $\widehat{\tau}_{\text{srd}}$
 2. The local linear estimator with bias-correction $\widehat{\tau}_{\text{srd}}^{\text{rbc}} = \widehat{\tau}_{\text{srd}} - \widehat{\text{bias}}$
- Two standard errors
 1. Standard SE $\widehat{\sigma}^2$
 2. “Robust” SE: accounts for uncertainty in bias estimation $\widehat{\sigma}_{\text{robust}}^2$
- We report the “Robust” estimate: $\widehat{\tau}_{\text{BC}}$ with $\widehat{\sigma}_{\text{robust}}^2$

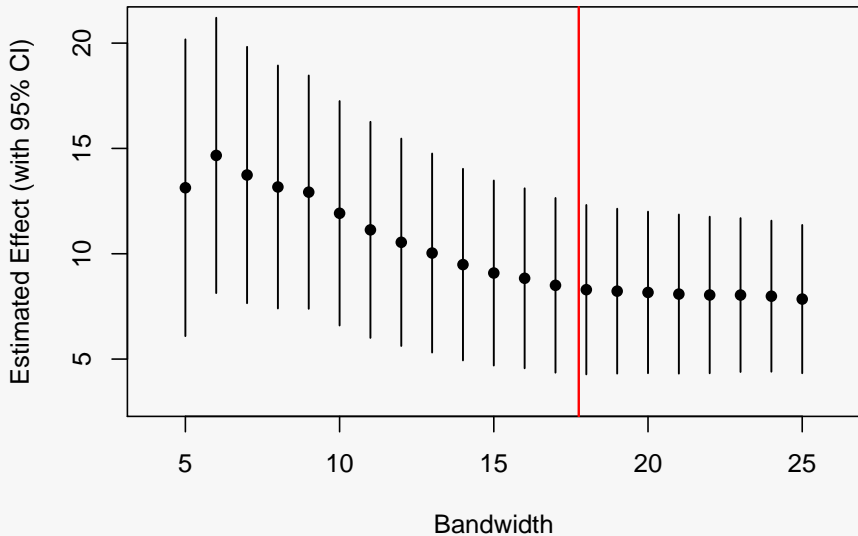
Estimated effect along different bandwidths

- We want to understand how results change along bandwidth

```
# fit local linear regression with bandwidth bws[b]
bws <- seq(5, 25, by = 1); fits <- list()
for (b in 1:length(bws)) {
  fits[[b]] <- rdrobust(y = vote, x = margin, h = bws[b],
    p = 1, kernel = "tri")
}

# summarize result (use "robust")
plot(1, 1, type = 'n', xlim = c(4, 26), ylim = c(3, 21),
  xlab = 'Bandwidth', ylab = 'Estimated Effect (with 95% CI)')
for (b in 1:length(bws)) {
  points(bws[b], fits[[b]]$coef[3], pch = 16)
  lines(c(bws[b], bws[b]), fits[[b]]$ci[3,], lwd = 1.2)
}
abline(v = fit$bws[1,1], col = 'red', lwd = 1.5)
```

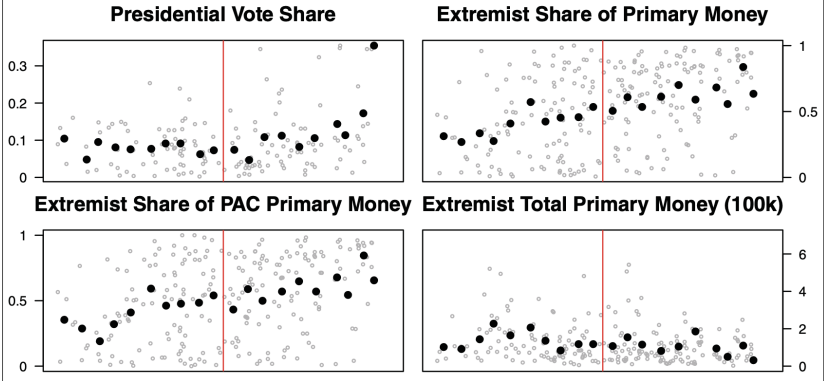
Result



Diagnostics: no-sorting?

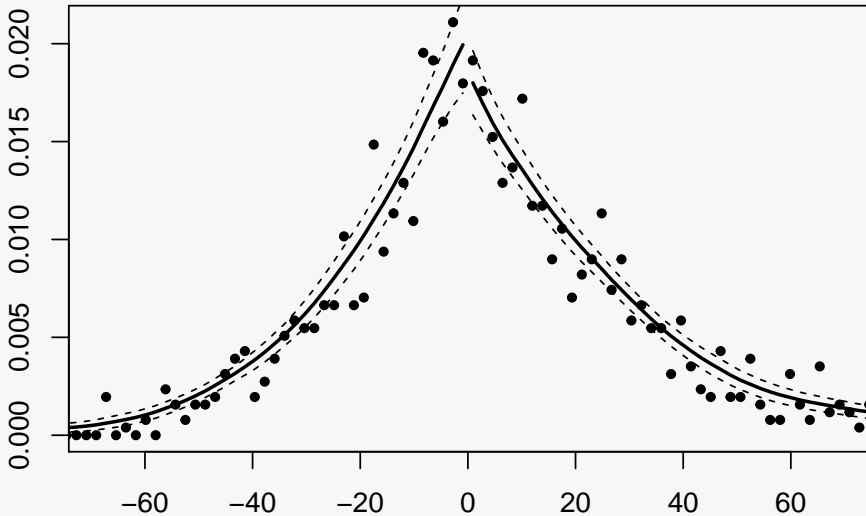
- Using pre-treatment covariates

FIGURE A.2. Graphical Balance Tests



Diagnostics: no-sorting?

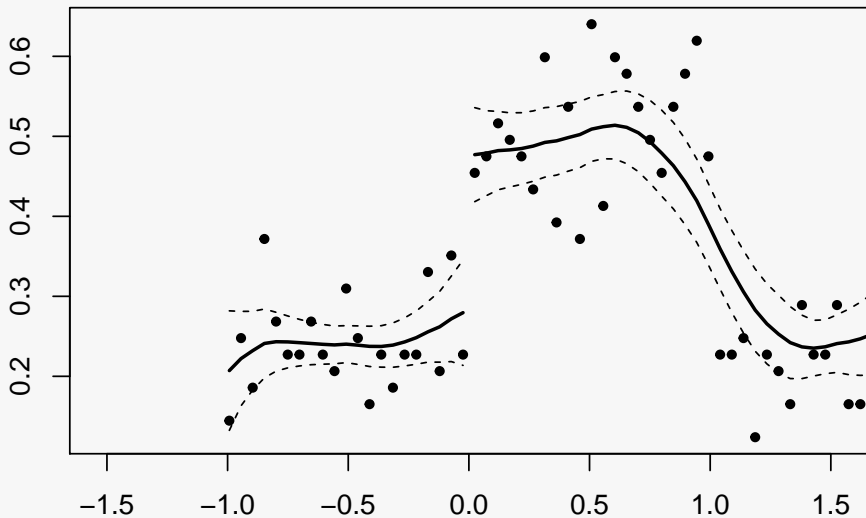
- McCrary test



[1] 0.3897849

Diagnostics: no-sorting?

- Example of discontinuity



[1] 0.02900635