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Gov 2002 (Harvard)

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- How should we summarize the distribution of causal effects?

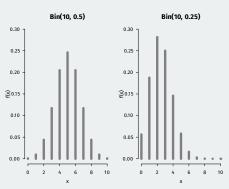
1/ Definition of Expectation

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 - · but we'll use our sample to learn about them

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- We'll use this intuition to create an average/mean for r.v.s.

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The **expected value** (or **expectation** or **mean**) of a discrete r.v. X with possible values, $x_1, x_2, ...$ is

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 - · Converse isn't true!

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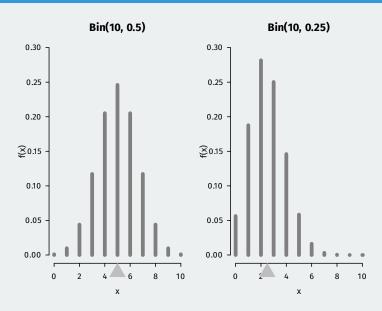
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Expectation as balancing point



2/ Linearity of Expectations

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 - $\mathbb{E}[XY] \neq \mathbb{E}[X]\mathbb{E}[Y]$ unless X and Y are independent.

Expectation of a binomial

• Let $X \sim \text{Bin}(n, p)$, what's $\mathbb{E}[X]$? Could just plug in formula:

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· Use linearity:

$$\mathbb{E}[X] = \mathbb{E}[X_1 + \dots + X_n] = \mathbb{E}[X_1] + \dots + \mathbb{E}[X_n] = np$$

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 Intuition: on average, the sample mean is equal to the population mean.

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 - If $X \ge Y$ with probability 1, then $\mathbb{E}(X) \ge \mathbb{E}(Y)$.

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 - Risk avoidance/concave utility $U = Y^{1/2} \leadsto \mathbb{E}[U(Y)] \approx 2.41$

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• Often, both of these are assumed away by assuming $\mathbb{E}[|X|] < \infty$ which implies $\mathbb{E}[X]$ exists and is finite.

3/ Indicator Variables

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• Use the fact that $\mathbb{I}(A_1\cup\cdots\cup A_n)\leq \mathbb{I}(A_1)+\cdots+\mathbb{I}(A_n)$ and then take expectations.

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- Use indicators! $I_j=1$ if jth condition is empty. So $I_1+\cdots+I_k$ is the number of empty conditions.

```
\begin{split} \mathbb{E}[I_j] &= \mathbb{P}(\mathsf{cond}\,j\,\mathsf{empty}) \\ &= \mathbb{P}(\{\mathsf{unit}\,\mathsf{1}\,\mathsf{not}\,\mathsf{in}\,\mathsf{cond}\,j\}\cap\cdots\cap\{\mathsf{unit}\,n\,\mathsf{not}\,\mathsf{in}\,\mathsf{cond}\,j\}) \\ &= \mathbb{P}(\{\mathsf{unit}\,\mathsf{1}\,\mathsf{not}\,\mathsf{in}\,\mathsf{cond}\,j\})\cdots\mathbb{P}(\{\mathsf{unit}\,n\,\mathsf{not}\,\mathsf{in}\,\mathsf{cond}\,j\}) \\ &= \left(1 - \frac{1}{k}\right)^n \end{split}
```

Using indicators to find expectations

- Suppose we are assigning n units to k treatments and all possibilities equally likely. What is the expected number of treatment conditions without any units?
- Use indicators! $I_j=1$ if jth condition is empty. So $I_1+\cdots+I_k$ is the number of empty conditions.

$$\begin{split} \mathbb{E}[I_j] &= \mathbb{P}(\mathsf{cond}\,j\,\mathsf{empty}) \\ &= \mathbb{P}(\{\mathsf{unit}\,1\,\mathsf{not}\,\mathsf{in}\,\mathsf{cond}\,j\}\cap \dots \cap \{\mathsf{unit}\,n\,\mathsf{not}\,\mathsf{in}\,\mathsf{cond}\,j\}) \\ &= \mathbb{P}(\{\mathsf{unit}\,1\,\mathsf{not}\,\mathsf{in}\,\mathsf{cond}\,j\}) \dots \mathbb{P}(\{\mathsf{unit}\,n\,\mathsf{not}\,\mathsf{in}\,\mathsf{cond}\,j\}) \\ &= \left(1 - \frac{1}{k}\right)^n \end{split}$$

• Thus, we have $\mathbb{E}\left[\sum_{i}I_{j}\right]=k(1-1/k)^{n}.$

4/ Variance

$$\mathbb{V}[X] = \mathbb{E}[(X - \mathbb{E}[X])^2]$$

• The **variance** measures the spread of the distribution:

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Useful equivalent representation of the variance:

$$\mathbb{V}[X] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$$

LOTUS

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LOTUS

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Defintion

The **Law of the Unconscious Statistician**, or LOTUS, states that if g(X) is a function of a discrete random variable, then

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The **Law of the Unconscious Statistician**, or LOTUS, states that if g(X) is a function of a discrete random variable, then

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• Example: $\mathbb{E}[X^2]$ where $X \sim \text{Bin}(n, p)$.

$$\begin{split} \mathbb{E}[X] &= \sum_{k=0}^n k \binom{n}{k} p^k (1-p)^{n-k} \\ \mathbb{E}[X^2] &= \sum_{k=0}^n k^2 \binom{n}{k} p^k (1-p)^{n-k} \end{split}$$

· Use LOTUS to calculate the variance for a discrete r.v.:

$$\mathbb{V}[X] = \sum_{j=1}^{k} (x_j - \mathbb{E}[X])^2 \mathbb{P}(X = x_j)$$

X	
0	1/8
1	3/8
2	3/8
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$$egin{array}{c|ccccc} x & p_X(x) & x - \mathbb{E}[X] \\ \hline 0 & 1/8 & -1.5 \\ 1 & 3/8 & -0.5 \\ 2 & 3/8 & 0.5 \\ 3 & 1/8 & 1.5 \\ \hline \end{array}$$

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x

$$p_X(x)$$
 $x - \mathbb{E}[X]$
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 1/8
 1.5
 2.25

$$V[X] = \sum_{j=1}^{k} (x_j - \mathbb{E}[X])^2 \rho_X(x_j)$$
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 - But this doesn't hold for dependent r.v.s
- 4. $V[X] \ge 0$ with equality holding only if X is a constant, $\mathbb{P}(X = b) = 1$.

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- Binomials are the sum of **independent** Bernoulli r.v.s so:

$$\mathbb{V}[X] = \mathbb{V}[X_1 + \dots + X_n] = \mathbb{V}[X_1] + \dots + \mathbb{V}[X_n] = np(1-p)$$

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 i!
 - We don't know what distribution it takes though!

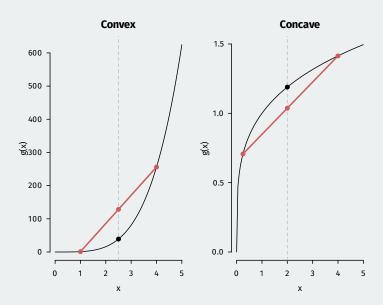
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- Remember that $\mathbb{E}[a+bX]=a+b\mathbb{E}[X]$ is linear, but $\mathbb{E}[g(X)]\neq g(\mathbb{E}[X])$ for nonlinear functions.
- Can we relate those? Yes for **convex** and **concave** functions.

Concave and convex



Jensen's inequality

Let X be a r.v. Then, we have

$$\mathbb{E}[g(X)] \ge g(\mathbb{E}[X])$$
 if g is convex $\mathbb{E}[g(X)] \le g(\mathbb{E}[X])$ if g is concave

with equality only holding if g is linear.

· Makes proving variance positive simple.

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 - $\mathbb{E}[1/X] \geq 1/\mathbb{E}[X]$
 - $\mathbb{E}[\log(X)] \leq \log(\mathbb{E}[X])$

6/ Poisson Distribution

Poisson

Definition

An r.v. X has the **Poisson distribution** with parameter $\lambda > 0$, written $X \sim \text{Pois}(\lambda)$ if the p.m.f. of X is:

$$\mathbb{P}(X = k) = \frac{e^{-\lambda} \lambda^k}{k!}, \qquad k = 0, 1, 2, ...$$

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- One more discrete distribution is very popular, especially for counts.
 - · Number of contributions a candidate for office receives in a day.
- Key calculus fact that makes this a valid p.m.f.: $\sum_{k=0}^{\infty} \lambda^k/k! = \mathrm{e}^{\lambda}$.

Poisson properties

• A Poisson r.v. $X \sim \text{Pois}(\lambda)$ has an unusual property:

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• If $X \sim Bin(n, p)$ with n large and p small, then X is approx Pois(np).